MARUDHAR KESARI JAIN COLLEGE FOR WOMEN, VANIYAMBADI

PG and Research Department of Mathematics

II B.Sc. Mathematics- Semester - III

E-Notes (Study Material)

Elective: Mathematical Statistics I

Code:23UEMA33

Unit:4: Poisson distribution: Properties, Moments of Poisson distribution– Geometric distribution:

Moment generating function of Geometric distribution.

Learning Objectives: To relate the statistical distributions with the real life situations.

Course Outcome: illustrate the theory of random variables, distribution functions and probability

distributions with suitable.

Overview: Poisson distribution, Geometric Distribution, Moment generating function

Unit 15 Poisson Distribution.

Condition.

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A Random variable × is said to be Poisson Distribution It it assumes only Non regative values and its probability mass function is given by mass function is given $-\lambda_{\lambda} x$; $x = 0, 1, 2, ..., \lambda = 0$ $P(x, \lambda) = P(x = 2c) = \begin{bmatrix} -\lambda_{\lambda} x \\ -\lambda_{\lambda} x \\ -\lambda_{\lambda} z \end{bmatrix}$; $x = 0, 1, 2, ..., \lambda = 0$ $D(x, \lambda) = P(x = 2c) = \begin{bmatrix} -\lambda_{\lambda} x \\ -\lambda_{\lambda} z \\ -\lambda_{\lambda} z \end{bmatrix}$; $x = 0, 1, 2, ..., \lambda = 0$

A Roachering Variable X, is Stud Where, 7 is known as parameter of the

distribution

Example : $x \sim p(x)$ denotes that x is a poisson variate with parameter ?

Note: $1 \stackrel{\infty}{\underset{x=0}{x=0}{\underset{x=0}{x=0}{\underset{x=0}{x=0}$ $=e^{-\lambda}e^{\lambda}=e^{0}=1$ $\sum \frac{\lambda^{2}}{\lambda} = e^{\lambda}$

Conditions for poisson Distribution.

1. NO of death, such as heart bettach by clue to snake bite, etc...

2. No of suicide in a particular city.

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3. No. of printing, in each page of the

book.

4. NO. OF car crossing for a minute during the busy of hours of a day. Moments OF poisson Distributions.

First four moments about origin of poisson Distributions are attained as follows.

 $\mu'_{i} = E(x) = \sum_{x=1}^{\infty} \alpha P(\alpha)$ $= \sum_{x=0}^{\infty} x \cdot \frac{e^{-\lambda_n}}{x!}$ $= \underset{x=0}{\overset{e}{\overset{n}}} x \dots \underbrace{\overset{e}{\overset{n}}}_{x(x-i)!}$ x-1 $= \sum_{x=0}^{\infty} \frac{e^{-\lambda_x^2} x^{1-1}}{(x-1)!} = \lambda e^{-\lambda} \sum_{x=0}^{\infty} \frac{\lambda}{(x-1)!}$ $= \partial e^{\partial} \left[\frac{\partial^{-1}}{\partial D} + \frac{\partial o}{\partial D} + \frac{\partial^{+1}}{\partial D} + \frac{\partial^{+1}}{\partial$ $= \lambda e^{\lambda} e^{\lambda}$ $\mu_{i} = \lambda$ $\mu_{i} = \lambda$ $\mu_{i} = E(x^{2}) = \bigotimes_{\substack{x=0 \\ x=0}}^{\infty} x^{2} p(x) = \bigotimes_{\substack{x=0 \\ x=0}}^{\infty} \frac{1}{2} p(x)$ $= \bigotimes_{\substack{x=0 \\ x=0}}^{\infty} x^{2} \frac{e^{\lambda} \lambda^{x}}{2} = \bigotimes_{\substack{x=0 \\ x=0}}^{\infty} (x^{2} + x - x) \frac{e^{\lambda} \lambda^{x}}{2}$ $= \sum_{x=0}^{\infty} [x(x-i)+x] e^{-\lambda} x^{x}$ $= \underbrace{\underbrace{\underbrace{\underbrace{\underbrace{x}}}_{x=0}}_{x=0} \times (x-i) \underbrace{\underbrace{e^{-\lambda}x}}_{x=1} + \underbrace{\underbrace{\underbrace{x}}_{x=0}}_{x=0} \underbrace{\underbrace{e^{-\lambda}x}}_{x=1}$ $= \underbrace{\underbrace{3}}_{x=0} x(x-i) \underbrace{e^{-\lambda} \lambda^{x}}_{x(x-i)i} + \lambda$

$$= \sum_{x=0}^{\infty} (x-i) \frac{e^{-\lambda} \lambda^{x}}{(x-i)(x-2)!} + \lambda$$

$$= \sum_{x=2}^{\infty} \frac{e^{-\lambda} \lambda^{x}}{(x-2)!} + \lambda = e^{-\lambda} \left[\sum_{x=0}^{\infty} \frac{\lambda^{x}}{(x-2)!} \right]_{\mu}$$

$$= e^{-\lambda} \left[\frac{\lambda^{2}}{0!} + \frac{\lambda^{3}}{1!} + \dots \right]_{\mu} + \lambda = e^{-\lambda^{2}} \left[e^{\lambda} \right]_{\mu} + \lambda$$

$$\mu_{3}' = \lambda^{3} + 3\lambda^{2} + \lambda$$

$$\mu_{4}' = \lambda^{4} + 6\lambda^{3} + 7\lambda^{2} + \lambda$$
(entral moments:

$$\mu_{1} = \mu_{1}' - \mu_{1}' = \lambda - \lambda = 0$$

$$\mu_{2} = \mu_{2}' - (\mu_{1}')^{2} = \lambda^{2} + \lambda - \lambda^{2} = \lambda$$

$$\mu_{3} = \mu_{3}' - 3\mu_{1}' \mu_{2}' + 2\mu_{1}'^{3}$$

$$= \lambda^{3} + 3\lambda^{2} + \lambda - 3(\lambda)(\lambda^{2} + \lambda) + 2\lambda^{3}$$

$$= \lambda^{3} + 3\lambda^{2} + \lambda - 3\lambda^{3} - 3\lambda^{2} + 2\lambda^{3}$$

$$\mu_{3} = \lambda$$

$$\mu_{4} = \mu_{4} - 4\mu_{3}'\mu_{1}' + 6\mu_{2}'\mu_{1}'^{2} - 3\mu_{1}'^{4}$$

$$= \lambda^{4} + 6\lambda^{3} + 7\lambda^{2} - 4(\lambda^{3} + 3\lambda^{2} + \lambda)\lambda + 6(\lambda^{2} + \lambda)\lambda^{2} - 3\lambda^{4}$$

$$= \lambda^{4} + 6\lambda^{3} + 7\lambda^{2} + \lambda + \lambda - 4\lambda^{4} - 12\lambda^{3}$$

$$-4\lambda^{2} + 6\lambda^{4} + 6\lambda^{3} - 3\lambda^{4}$$

$$\mu_{4} = 3\lambda^{2} + \lambda$$

Skewnees

$$B_1 = \frac{\mu_3^2}{M_2^3} = \frac{\lambda^2}{\lambda^3} = \frac{1}{\lambda}$$

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$$\beta_{2} = \frac{\mu_{1}}{\mu_{2}} = \frac{2\lambda^{2}+\lambda}{\lambda^{2}} = \frac{2\lambda^{2}}{\lambda^{2}} + \frac{\lambda}{\lambda^{2}}$$

$$= 3 + \sqrt{\lambda}$$
Moment generating function of Risson
distribution.

$$M_{x}(t) = E(e^{t \times t}) = \sum_{x=0}^{\infty} e^{t \times t} p(x)$$

$$= \sum_{x=0}^{\infty} e^{t \times e^{t} \lambda x}$$

$$M_{x}(t) = e^{\lambda} \sum_{x=0}^{\infty} e^{t \times \lambda^{2}}$$

$$= e^{\lambda} \left[1 + \frac{e^{\lambda}}{11} + \frac{e^{2t}\lambda^{2}}{2!} + \dots \right]$$

$$= e^{\lambda} \left[e^{t} - \lambda \right] = e^{\lambda} \left[e^{t} \lambda \right]$$

$$M_{x}(t) = e^{\lambda} \left[e^{t} - 1 \right] = e^{\lambda} \left[e^{t} \lambda \right]$$

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$$M_{x}(t) = e^{\lambda} \left[e^{\lambda} + 1 \right]$$

$$M_{x}(t) = e^{$$

Proof

$$\begin{aligned} \bigcap_{i=1}^{n} x_i(t) &= M(x_i + 2(2 + \dots + 2n)(t)) \\ &= Mx_i(t) \cdot Mx_2(t) \cdot \dots \cdot Mx_n(t) \\ &= Mx_i(t) \cdot Mx_2(t) \cdot \dots \cdot Mx_n(t) \\ &= e^{\lambda_i(e^{t-1})} \cdot \lambda_2(e^{t-1}) \cdot \dots \cdot e^{\lambda_n(e^{t-1})} \\ &= e^{(e^{t-1})} \left(\sum_{i=1}^{n} \lambda_i \right) \quad M_x(t) = e^{\lambda(e^{t-1})} \end{aligned}$$

Which is the moment generating function of poisson variate with parameter d, dz, dy

Hence by uniqueness theorem of moment generating function, $\leq X_i$ is also a poisson variate.

Remark

The difference of two independent Poisson variate is not a poisson variate $M_{x_1} - x_2(t) = M_{x_1}(t)$. $H_{x_2}(t)$ $= (M_{x_1} \cdot M_{x_2})t = M_{x_1}(t) \cdot M_{x_2}(t)$ $= e^{\lambda_1(e^t - 1)}e^{\lambda_2(e^t - 1)}$ $= e^{\lambda_1(e^t - 1)}t^{\lambda_2}(e^{-t} - 1)$ which cannot be putting the form $e^{\lambda}(e^t - 1)$

:. x, -X2 Is not a poisson variate.

Recurrance Formula for the Probability of poisson distribution (or) fitting of poisson distribution.

The poisson distributing with A is depind by $P(x) = \frac{e^{\lambda} \lambda^{x}}{\infty!}$ $(2) = (1) = \frac{e^{-\lambda} A^{2+1}}{(x+0)!} = \frac{e^{-\lambda} A^{2+1}}{(x+0)!} \times \frac{2!}{e^{-\lambda} A^{2}}$ $(2) = (1) = \frac{P(x+1)}{P(x)!} = \frac{e^{-\lambda} A^{2+1}}{(x+1)!} \times \frac{2!}{e^{-\lambda} A^{2}}$ $\frac{P(x+i)}{P(x+i)} = \frac{\lambda}{(x+i)!} \frac{x}{(x+i)!} \frac{x}{(x+i)!} \frac{x}{(x+i)!} \frac{x}{(x+i)!} \frac{x}{(x+i)!}$ $P(x+i) = \frac{\partial}{\partial x+i} P(x)$ which is the required recurrance formula only probability we need the calculate plo) is given by $P(\alpha) = \frac{\epsilon^{\lambda} \partial^{\alpha}}{\partial c^{\lambda}}$ (1) set (1) $P(o) = \overline{e^{\lambda}} \overline{\lambda^{o}} = \overline{e^{-\lambda}}$ Where it is estimated using given data The another Probability is p(1), P(2)... i.e. mean = A be easily obtain as follows P(x+1) x===) P(1) = - P[0)

 $P(1) = \pi e^{-\pi}$

$$P(x+i)_{x=1} \Rightarrow P(2) = \frac{\lambda}{2} P(i)$$

$$P(2) = \frac{\lambda}{2} \lambda e^{\lambda}$$

$$P(3) = \frac{\lambda^{2}}{2} e^{\lambda}$$

$$P(3) = \frac{\lambda^$$

when it is not an integer let us Suppose that 's' is the integral part of the it

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$$\frac{p(x)}{p(x-i)} \text{ for } x = 1, 2, \dots S$$

$$\frac{p(x)}{p(x-i)} \ge 1 \text{ for } x = S + 1, S + 2, \dots S + n$$

$$p(x-i) \ge p(x-i) = p(x-i) \dots p(s) > p(S-i)$$

p(s+1) ~ P(s), p(s+2) < p(s+1) ... 00 =>p(0) < p(1) < p(2) ... p(S-D < P(S) =) P(s) > P(s+1) > P(s+2) ... combaining above expression we get p(0) < p() < p(2) ... < p(s) > p(s+1) > p(s+2) In this case the distribution is unique model and the integral part of his unique model value. case (ii) A is an integer, A=K $\frac{P(x)}{P(x-0)} = \begin{cases} >1 & \text{for } x = Y_{k}(\text{or}) \ 0, 1, 2, \dots \\ =1 & \text{for } \mathcal{X} = k \\ <1 & \text{for } \mathcal{X} = k + 1, k + 2 \end{cases}$ =>p(i)>p(o), p(2)>p(i), ... p(K-1)>p(K-2) P(6) & P(1), P(1) & P(2) ... P(1<-2) < P(K-1) $\frac{P(x)}{P(x)} = 1$ when $x = k+1, k+2, \dots$ $\frac{P(x)}{P(x)} = 1$ $P(x-1) = 1 \qquad \frac{P(x-1)}{P(x-1)} < 1$ When $\mathcal{D}(=K \frac{P(K)}{P(K-1)} = 1$ $\frac{P(K+1)}{P(K+1-1)} \leq 1$ $\frac{P(k+1)}{P(k+1)} < 1 \dots \infty$ $\frac{P(k+1)}{P(k)} < 1$ $P(k+1) < p(k), p(k+2) < p(k+1) \dots \infty$ combairing the above expression, we get P(6)<P(1) < P(2) ... < P(K-D= P(K)>P(K+1)> FLK+2)1. Therefore we have two masurmum Value P(K-1) and p(K) the distribution is

bimodal are K-1 & K

i.e., which is N-192

K= >

Problems:

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A car hive firm has two cars which its fires out day by day. Then the number of demand on each day is distributed has poisson variate with mean 1.5, calculate the preposition of plays on which
i) Neither car is used
i) Some demand is refused.

Here the random variable X which denotes the no. of demands for a car or any day follows poisson distribution. with mean $\lambda = 1.5$

The propertion of days which there is a demands for a car is given by A=15

 $P(x) = \frac{e^{-\lambda_{1}x}}{x!} = e^{\frac{1.5}{(1.5)^{\alpha}}} x_{=0,1,2,...}$ i) Neither car is used Propertion of days on which

heither car is used.

$$P(x=0) = e^{\frac{1}{5}} (1.5)^{\circ} = e^{1.5}$$

i) Some demand is rejused
Propertion of days on which
Some demand is refused is

$$P(x>2) = 1 - P(x \le 2)$$

 $= 1 - \left[P(x=0) + P(x=1) + P(x=2)\right]$
 $= 1 - e^{1.5} \left[1 + 1.5 + \left(\frac{1.5}{2}\right)^2\right]$
 $P(x>2) = 0.1912$
2) A manufacture of cotton pipes knows
 $= 1 - e^{1.5} \left[1 + 1.5 + \left(\frac{1.5}{2}\right)^2\right]$

that 5% off is product it is defective Sales, cotton ping in boxes are 100 and guarentee quality not more than 10 ping will be defective, what is the approximate probability that a box will fail to meet the gurentee quality? Griven: h=100 Let, P= Probability of defective items 5%.

 $P = \frac{5}{100} = 0.05$ A = mean = NO.05 débertive pins in a box of 100

$$\lambda = np$$

$$\lambda = 100 \times 0.05 = 5$$

$$P(x) = \frac{e^{-n} \lambda^{x}}{2^{c}!} = \frac{e^{-5} (5)^{x}}{2^{c}!}, x = 0, 1, 2, \dots$$
Probability of guass will fair to meet
the aircorb ac angulation is

= 100-10 = 90 > 10

 $P(x>10) = 1 - P(x \le 10)$ $= 1 - [P(x=0) + P(x=1) + P(x=2) + P(x=3)]_{1}$ p(x = 4) + p(x = 5) + p(x = 6) +p(x = -1) + p(x = 8) + p(x = 9) + p(x = 9)

한번에 가지 않는 것 같아. 아이들 것이 같아.

이 가는 것은 것을 가지만 하는 것이 없는 것이 같은 것이 없다.

an ward where don't filling a statistic

3) An isnsurrance company insures 4000 People against loss of both eyes in a car accident. Based on previous data, the rates were computed on the assumption that on the average 10 persons in 1,00,000 will have can accident each year that result in this type of Injury what is the probability that more than 3 of the insured with collect on their policy in a given year?

In usual Notations, given,

n = 4000 and P = probability of lossof both eyes in a car accident $= <math>\frac{10}{100000} = 0.0001$

Since P is very small and the given distribution
may approximate the given distribution
by poisson distribution thus the parameter

$$\lambda$$
 of the poisson distribution is,
 $\lambda = n P = 4000 \times 0.0001 = 0.4$
Let the random variable x denotes
number of car accidents in batch of
 $4,000$ People then by Poisson probability
law,
 $p(x = x) = e^{-\lambda} \frac{\lambda^{x}}{x_{1}} = e^{-0.4} (0.4)^{x} (0.4)^{x}$
 $p(x = x) = e^{-\lambda} \frac{\lambda^{x}}{x_{1}} = e^{-0.4} (0.4)^{x} (0.4)^{x}$
Here, the required probability that more
than 3 off the insured will what more
their Policy is given by,
 $p(x = 2) + p(x = 3)$
 $= 1 - e^{-0.4} [(0.4)^{2} + (0.4) + (0.4)^{2} + \frac{21}{2!} + \frac{(0.4)^{3}}{2!} \frac{2}{1!}$
 $= 1 - 0.6703 (1 + 0.4] + \frac{3!}{3!} \int_{0.08 + 0.0107}^{0.08 + 0.0107}$

=0.0008

») A Manufacturer, who produces medicine bottles, find that 0.1%. Of the bottles are defective. The bottles are packed in boxes containing 500 bottles. A drug manufac turer buys 100 boxes brom the producer of bottles using poisson destribution. Find how many boxes will contain. i) No defective and

ii) at least two defective.

In the usual notations, given : N = 100, N = 500, p = probability of a defective bottle = 0.001,

A=np=500x0.001=0.5

Let the random variable × denote the number of defective bottles in a box of 500. Then by poisson probability law, the probability of x defective bottles in a box is given by,

 $P(x=oc) = e^{-0.5} (0.5)^{\alpha} = 0.6065x(0.5)^{\alpha}$

56=0,1,2,...

Hence in a consignment of 100 boxes, the frequency (number) of boxes

containing x defective bottles is, Two values $f(x) = NP(x=x) = 100 \times 0.6065 \times (0.5)^{\alpha}$

i) Number of boxes containing no defective bottles,

 $= 100 \times P(x=0) = 100 \times 0.6065$

11) NO. of Boxes containing at least two defective bottles, 1 - [P(x=0) + P(x=1)]so 0,1 = 100 [P(x=2)] = 100 [1 - P(x=0) - P(x=1)]

=100 (1-0.6065-0.6065 x0.5)

=100 × 0.09025 = 9



Statistics.

H) In a book of 520 pages, 390 typo-graphical errors occur. Assuming Poisson law for the number of errors per page, find the probability that a random sample of 5 pages will contain no error.

The average no. of typo-graphical errors per page in the book is given by

$$\lambda = \frac{390}{520} = 0.75$$

Hence, By Poisson probability law, the probability of x errors per page is given by

$$P(x=x) = e^{-\lambda x}, x = 0, 1, 2, ...$$

The required probability that a random Sample of 5 pages will contain no error is

$$\begin{bmatrix} P(x=0) \end{bmatrix}^{5} = \begin{bmatrix} e^{-0.75} (0.75)^{\circ} \\ 0! \end{bmatrix} = (e^{-0.75})^{5}$$
$$= e^{-3.75}$$
$$\begin{bmatrix} P(x=0) \end{bmatrix}^{5} = 0.0235$$

parameter is a

The frequency function is given by f(sc) = NP(sc) = NP[sc=sc]

= $N \cdot \frac{e^{\lambda} x}{x_1}; x = 0, 1, 2, \dots \rightarrow 0$ put x=3 in 1 $f(3) = N \cdot \frac{e^{-\lambda_{1}}}{21}$ put x=4 in O $f(y) = N \cdot \frac{e^{-\lambda y}}{1 \cdot 1}$ $f(3) = \frac{2}{3}f(4)$ $N \cdot \frac{e^{-\lambda}}{2!} = N \cdot \frac{2}{3} \cdot \frac{e^{-\lambda}}{1!}$ 13: - 2 2 $\frac{1}{6}\frac{3}{2}=\frac{3}{21}$ 7= 6 Mean of the poisson distribution is equal to? 1=6

Standard deviation in poisson distribution is S. D=56

6) Suppose That the number of telephone calls coming int 0 a telephone exchange between 10 a.m and 11 am Say, x, is a random variable with poisson distribution with parameter <u>two</u>. Similarly The number of calls arriving between 11 am and 12 noon Say, X2 thas a poisson distribution with parameter 6. It x, and X2 are independent, what is the probability that more than 5 calls come in between 10 am and 12 noon.

briven : X1~p(2) and X2~p(6) Let $X = X_1 + X_2$ By the additive property of poisson distribution, X is also a poisson variate with parameter λ. i.e) $\lambda = 2+6$ 2=8 Hence, The probability of 24 calls in between 10 am and 12 noon is given by. $P(x = \underline{x}) = \underline{e^{-n}} \underline{\lambda}^{x}$ $= \frac{e^{-8}8^{x}}{2}; x=0, 1, 2...$ The probability that more than 5 calls in between 10 am and 12 noon. $P(x>5) = 1 - P(x \le 5)$ $= 1 - \left(P(x=0) + P(x=1) + P(x=2) + P(x=3) + P$ P(x = 4) + P(x = 5) $= 1 - \left[\frac{e^{-8}8}{1} + \frac{e^{-8}8}{1} + \frac{e^{-8}8}{2} + \frac{e^{-8}8}{2} + \frac{e^{-8}8}{2} + \frac{e^{-8}8}{4} + \frac{e$ $= 1 - \left[e^{-\frac{8}{5}} \right] + 8 + 32 + 85 \cdot 3 + 170 \cdot 67 + 2730 \right]$ =1-[=8 (569.97)] =1-0.1912036 P(X75) = 0.80879 7) A poisson distribution It x is a poisson variate such that P(x=2) = qp(x = 4) + qop(x = 6). Find ii) Mean (ii) B, the coefficient of skewness i) y of X If x is a poisson variate with parameter $P(x=2) = 9 P(x=4) + 90 P(x=6) \rightarrow \bigcirc$ 2.

$$P(x=x) = \frac{e^{-\lambda_{\lambda}x^{2}}}{x!}; x=0,1,2,...$$

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From (i)

$$P(x=2) = q P(x=u) + 90 P(x=b)$$

$$\frac{e^{\lambda} \lambda^{2}}{2!} = q \frac{e^{\lambda} \lambda^{4}}{4!} + 90 \frac{e^{\lambda} \lambda^{6}}{6!}$$

$$\frac{e^{\lambda} \lambda^{2}}{2!} = e^{-\lambda} \left[\frac{q \lambda^{4}}{2u} + \frac{90 \lambda^{6}}{720}\right]$$

$$\frac{e^{\lambda} \lambda^{2}}{2!} = e^{-\lambda} \left[\frac{q \lambda^{4}}{2u} + \frac{90 \lambda^{6}}{720}\right]$$

$$\frac{e^{\lambda} \lambda^{2}}{2!} = e^{-\lambda} \lambda^{2} \left[3 \lambda^{2} + \lambda^{4}\right]$$

$$1 = \frac{1}{4!} (3 \lambda^{2} + \lambda^{4})$$

$$\lambda^{4} + 3 \lambda^{2} + = 0$$

$$put \lambda^{2} = y,$$

$$y^{2} + 3y - 4 = 0$$

$$put \lambda^{2} = y,$$

$$y^{2} + 3y - 4 = 0$$

$$y + 4 = 0$$

$$y + 4 = 0$$

$$y - 1 = 0$$

$$y + 4 = 0$$

$$y - 1 = 0$$

$$y + 4 = 0$$

$$y - 1 = 0$$

$$y + 4 = 0$$

$$y - 1 = 0$$

$$y + 4 = 0$$

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$$y + 4 = 0$$

$$y + 4 = 0$$

$$y - 1 = 0$$

$$y + 4 = 0$$

$$y + 4 = 0$$

$$y + 4 = 0$$

$$y + 1 = 0$$

$$y + 4 = 0$$

$$y$$

Let X~ p(A) and Y~P(M) We know that, $P(x=\alpha) = \frac{e^{\lambda} \partial^{\alpha}}{e^{\lambda}}$ From ()=) $e^{-\lambda_1} = e^{-\lambda_1^2}$ and $e^{-\mu_1^2} = e^{-\mu_1^3}$ $\lambda e^{-\lambda} = e^{-\lambda} \lambda^2$ and $\mu^2 e^{-\mu} = \mu^3 e^{-\mu} \rightarrow \bigcirc$ $\mu^2 e^{-\mu} = \mu^3 e^{-\mu}$ Solving ean (2) Nen= Nen 3 = H 27e2 = 2e2 µ=3 7=2 an gin bel press per parteret ar had a press and 9) Fit a Poisson distribution to the following data

Number of mistakes per page 0 1 2 3 4 Total Number of pages : 109 65 22 3 1 200 If the above distribution is approximated by a Poisson distribution, then the parameter of Poisson distribution is given by.

$$m = 0 \times 107 + 1 \times 105 + 2 \times 22 + 3 \times 3 + 4 \times 1$$

$$= 0.5$$

$$200$$
By poisson probability law, the frequency (number)
of page centaining r mistakes is given by
$$f(r) = N \cdot P(r) = 200 \times \frac{2^{-0.61} (0.61)^{2}}{\pi!}$$
Putting $r = 0, 1, 2, \dots$ we get the expected
brequencies of poisson distribution.
$$[A(50, P(0) = e^{-0.6} = (2.71828)^{-0.61} = \frac{1}{Antilog (0.61 \times log 2.7182g)}$$

$$= \frac{1}{Antilog (0.61 \times 0.4343)} = \frac{1}{Antilog (0.26492)}$$

$$= \frac{1}{1 \cdot 8441} = 0.5432.7 \times$$
Calculation for expected poisson brequencies.
No. of Histokes
$$0 \qquad f(0) = 200 \times e^{-0.61} \times (0.61)^{2} = 20.21 \qquad 20$$

$$3 \qquad f(3) = 200 \times e^{-0.61} \times (0.61)^{3} = 4.11 \qquad 24$$

$$4. \qquad f(4) = 200 \times e^{-0.61} \times (0.61)^{4} = 0.63 \qquad 21$$

the proof reader finds that there are, on the average 2 errors per 5 pages that makey pages would one expect to find with 0, 1, 2, 3 and 4 errors, in 1000 pages of the first print of the book? Let the random variable × denote The number of errors per page. Then the mean number of errors per page is given by $A = \frac{2}{5} = 0.4$ Using poisson probability law, probability of J errors per page is given by $P(x=x)=P(x)=\frac{e^{-1}h^{x}}{x!}=\frac{e^{-0.4}(0.4)^{x}}{x!}; x=0,1,2,...$ Expected number of pages with x errors per page in a book of N=1000 pages are. $f(x) = N P(x = x) = 1000 \times e^{-0.4} (0.4)^{2}, x = 91, 2...$ Using the recurrance. No. of errors probability Expected number of page f(x) = Np(x) PLOC) per page (x) 670.32670 p(0)= e = 0.6703 0 268.12 2268 $P(1) = \frac{0.4}{0+1} P(0) = 0.26812$ 1 53.624 2 54 P(2) = 0.4 P(1) = 0.0536242 P(3) = 0.4 p(2) = 0.71298 7.1298 273 P(4) = 0.4 P(3) = 0.0007129 0.71298 ~13+1 4

10) After correcting 50 pages of the proof of a book,

1) Fit a poisson distribution to the bollowing data which gives the number of doddens a sample of cloves seeds No. of deeds (x) 0 1 2 3 4 5 67 ŧ Observed frequency (+) 56 156 132 92 37 22 4 0 Mean = $\frac{1}{N} \leq f(x) = \frac{986}{500} = 1.972$ Taking the mean of the poisson distribution we want to fit, we get 2=1.972. $P(x) = \frac{e^{-\lambda} a^{x}}{-(1)}; x = 0, 1, 2, ...$ P(0)= e = e = 1.972 log, p(0)= - 1.972 log, e = -1.972 XO. 48424 =-0.856419 = T.143581 P(0) = Antilog (7.1436)=0.1392 Using the recurrence formula, the various probabilities, viz, p(D,p(2),... (an be easily calculated as shown in the tollowing table. $\frac{\lambda}{x+1}$ C p(x) Expected brequency 1.972 0.13920 0 NP(x) = 500.P(x)69.6000270 0.986 0.27455 ١ 137.2512 ~ 137 0.657 0.27006 135.32962135 2 3 0.493 0.17793 86.9566 289 0.394 0.10964 4 43.6556 244

5	0.328	0.03459	17.2966 ~17
6	0.281	0. 01137	5.684626
7	0.247	0.00320	1-60132 2
8	0.219	0.00078	0.394220

(reometric Distribution.

A random variable X is said to have a geometric distribution. If it assumes only non-negative values and its probability mass function is given by $P(x=x) = \int q^{x}p$; $x=0,1,2,..., 0 \le p \le 1; q=1-p$ 0; otherwise.

Moments of Geometric distribution Mean = $\mu_i' = \Xi x p(x = x)$ $= \underbrace{\sum_{x=1}^{\infty} DL}_{x=1} Pq_{x} = Pq_{x} \underbrace{\sum_{x=1}^{\infty} DLq_{x}}_{x=1} x^{-1}$ Mean = $Pq(1-q)^2 = \frac{q}{p}$ $V(x) = E(x^{2}) - [E(x)]^{2}$ $= E \left[\times (\times -i) \right] + E (\times) - \left[E(X) \right]^{2}$ $E[x(x-i)] = \overset{\infty}{\leq} x(x-i) P(x=x)$ $= \sum_{x=2}^{\infty} c(x-1) P q^{x}$ $= 2pq^{2} \sum_{x=2}^{\infty} \left[\frac{x(x-1)}{2x} q^{x-2} \right]$ $= 2pq_{1}^{2}(1-q_{1})^{-3} = \frac{2q_{1}^{2}}{p^{2}}$ $\therefore V(x) = \mu_{2} = \frac{2q_{1}^{2}}{p^{2}} + \frac{q_{1}}{p} - \frac{q_{1}^{2}}{p^{2}} = \frac{q_{1}^{2}}{p^{2}} + \frac{q_{2}}{p} = \frac{q_{1}}{p^{2}}$

Moment generating bunction of Geometric Distribution $M_{X(t)} = E[e^{tX}] = \underset{T=0}{\overset{\infty}{\underset{T=0}{\overset{}}} e^{tX}.q^{T}.P$ = $P \stackrel{2}{=} (e^{t}q)^{\chi} = P(1 - q_{e}t)^{-1} = \frac{P}{1 - q_{e}t}$ $M_{i}' = \begin{bmatrix} d \\ dt \end{bmatrix}_{k=0} = \begin{bmatrix} d \\ dt \end{bmatrix} (1 - qe^{k})^{-1} \end{bmatrix}_{k=0}$ $= Pq((1-q))^{2} = q/p$ $M_2' = \left[\frac{d^2}{dt^2} + M(t)\right]_{t=0} = \frac{9}{P} + \frac{29^2}{P^2}$ $H_2 = H_2' - H_1' = \frac{q_1}{p} + \frac{2q_1^2}{p^2} - \frac{q_1^2}{D^2}$ $= q^{2} + pq = q (q+p)=1$ $p^{2} = p^{2}$ Hence the mean and variance of geometric distribution-are p and q respectively. Remark !) Variance = $\frac{q}{p^2} = \frac{q}{p} \cdot \frac{1}{p} = \frac{Mean}{p} > mean$ Hence, For the Geometric distribution variance is greater than the mean. 2) The Probability generating function of the geometric distribution is obtained on replacing et by s P Px

$$(s) = \frac{1}{(1-9,s)}$$

Practice Questions:

2 Marks:

- 1. Define Moment Generating Function of poisson distribution
- 2. Find rth moment x about origin
- 3. Write the Characteristic function of binomial distribution.
- 4. Six coins are tossed 6400 times using the Poisson distribution, find the approximate probability of getting 6 heads r times.

5 Marks:

- 1. If X is said to be a continuous random variable with p.d.f. $f(x) = 6x(1-x), 0 \le x \le 1$, Check that f(x) is p.d.f.
- 2. Effect of change of origin and scale on moment generating function.
- 3. Explain additive property of Poisson distribution.
- 4. Derive the Mode of Poisson distribution.
- 5. A car hire firm has two cars which its fires out day by day.then the number of demand on each day is distributed has poisson variate with mean 1.5, calculate the preposition of days of which (i) Neither car is used (ii) some demand is refused.

10 Marks:

- 1. Derive Reproductive property of independent of Poisson distribution.
- 2. Explain fitting of Poisson distribution.
- 3. After orreting 50 pages of the proof of a book, the proof reader finds that there are on the average of 2 errors per 5 pages. How many pages would are expect to find out with 0,1,2,3 and 4 errors in 1000 pages of the first print of the book.

References:

S.C. Gupta and V.K. Kapoor, Fundamentals of Mathematical Statistics, Sultan Chand &Co, New Delhi, Reprint 2019.